

Lecture 6. Threshold functions of random graphs.

We have showed that surprisingly many tempting conjectures can be easily disproves by the probabilistic method and random graphs.

Question: For which $p = p_n$ is a property true for $G(n, p)$ w.h.p.?

Graph property: A graph property \mathcal{P} is a subset of all graphs.

monotone (increasing): $\forall G \in \mathcal{P}$, any supergraph (adding edge) $\in \mathcal{P}$.

non-trivial: \forall sufficiently large n , \exists n -vertex graph $G \notin \mathcal{P}$.

Examples: containing H (monotone, increasing), connected

planar (monotone, decreasing), containing vertex of deg 1 (nor).

Warmup: $\mathcal{P} = \{G : \text{triangle } K_3 \subseteq G\}$. consider $G \sim G(n, p_n)$.

Let $X = \# \text{ of } \Delta s$. $\mathbb{E}[X] = \binom{n}{3} p^3 \approx n^3 p^3$. If $p < \frac{1}{n}$, $\Pr[X \geq 1] = o(1)$ by Markov's inequality. If $p > \frac{1}{n}$, $\mathbb{E}[X] \rightarrow \infty$. Markov fails.

Chebyshov: $\Pr[|X - \mathbb{E}[X]| \geq t] \leq \frac{\text{Var}[X]}{t^2} \Rightarrow \Pr[X = 0] = o(1)$ if

$X \geq 0$, and $\text{Var}[X] = o(\mathbb{E}[X]^2)$. If $X = X_1 + \dots + X_n$ then $\text{Var}[X]$

$= \sum \mathbb{E}[X_i] + 2 \sum_{i < j} \text{Cov}[X_i, X_j]$, $\text{Cov}[X, Y] = \mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y]$.

$\text{Cov}[X_{T_1}, X_{T_2}] = \mathbb{E}[X_{T_1}X_{T_2}] - \mathbb{E}[X_{T_1}]\mathbb{E}[X_{T_2}] = P^{e(T_1 \cup T_2)} - P^{e(T_1) + e(T_2)}$

$$= \begin{cases} 0 & |\bar{T}_1 \cap \bar{T}_2| \leq 1 \\ p^5 - p^6 & |\bar{T}_1 \cap \bar{T}_2| = 2. \end{cases} \quad \text{Var}[\bar{X}_{\bar{T}}] = |\mathbb{E}[\bar{X}_{\bar{T}}^2] - \mathbb{E}[\bar{X}_{\bar{T}}]^2| = p^3 - p^6.$$

$$\Rightarrow \text{Var}[\bar{X}] = \sum_{\bar{T}} \text{Var}[\bar{X}_{\bar{T}}] + \sum_{\bar{T}_1 \neq \bar{T}_2} \text{Cov}[\bar{X}_{\bar{T}_1}, \bar{X}_{\bar{T}_2}] = \binom{n}{3} (p^3 - p^6) + \sum_{|\bar{T}_1 \cap \bar{T}_2|=2} (p^5 - p^6)$$

$$= \binom{n}{3} (p^3 - p^6) + \binom{n}{2} (n-2)(n-3) (p^5 - p^6) \lesssim n^3 p^3 + n^4 p^5. \text{ We hope}$$

$$\text{Var}[\bar{X}] = o(\mathbb{E}[\bar{X}]^2) \Leftrightarrow n^3 p^3 + n^4 p^5 = o(n^6 p^6) \Leftrightarrow p \gg 1/n.$$

Definition. (Threshold function). We say r_n is a threshold function

for some graph property \mathcal{P} . If $\Pr[G(n, p_n) \in \mathcal{P}] \rightarrow \begin{cases} 0 & r_n/p_n \rightarrow 0 \\ 1 & r_n/p_n \rightarrow \infty. \end{cases}$

Theorem. A threshold function for containing a K_3 is $1/n$.

Question. What about containing a K_4 ? (already known $n^{-2/3}$).

General setup. Suppose $X = X_1 + \dots + X_m$, X_i is indicator of event E_i .

Denote by $i \sim j$ if $i \neq j$ but E_i, E_j are not independent. (for variance.

we only need pairwise independence, but will later rely on a stronger

notion of dependency in Lovász local lemma). If $i \neq j$ and $i \not\sim j$

clearly $\text{Cov}[X_i, X_j] = 0$, otherwise $\text{Cov}[X_i, X_j] \leq \mathbb{E}[X_i X_j] = \Pr[E_i \wedge E_j]$

Thus $\text{Var}[X] = \sum_{i,j} \text{Cov}[X_i, X_j] \leq \mathbb{E}[\bar{X}] + \underbrace{\sum_{i \sim j} \Pr[\bar{E}_i \wedge \bar{E}_j]}_{\Delta} \triangleq \Delta.$

We hope $\text{Var}[\bar{X}] = o(\mathbb{E}[\bar{X}]^2)$. So $\Delta = o(\mathbb{E}[\bar{X}]^2)$ suffices. Furthermore.

$\sum_{i \sim j} \Pr[\bar{E}_i \wedge \bar{E}_j] = \sum_i \Pr[\bar{E}_i] \sum_{j \sim i} \Pr[\bar{E}_j | \bar{E}_i]$. In many symmetric cases.

$\sum_{j \sim i} \Pr[\bar{E}_j | \bar{E}_i]$ does not depend on i . Denote it by Δ^* . (or we may set $\Delta^* = \max_i \sum_{j \sim i} \Pr[\bar{E}_j | \bar{E}_i]$). $\Rightarrow \Delta = \sum_i \Pr[\bar{E}_i] \Delta^* = |\bar{E}[X]| \Delta^*$.

Lemma. If $|\bar{E}[X]| \rightarrow \infty$ and $\Delta^* = o(|\bar{E}[X]|)$, then $X > 0$ w.h.p.

Now consider the property of containing K_4 . $\forall S, \bar{E}_S = S$ forms a K_4 .

$\forall S, T$ of size 4. $S \sim T$ iff $|S \cap T| \geq 2$. There are two cases :

$$\textcircled{1} \quad |S \cap T| = 2. \quad \sum_T \Pr[\bar{E}_T | \bar{E}_S] \leq 6 \binom{n}{2} \Pr[\bar{E}_T | \bar{E}_S] = 6 \binom{n}{2} p^5 \approx n^2 p^5.$$

$$\textcircled{2} \quad |S \cap T| = 3. \quad \sum_T \Pr[\bar{E}_T | \bar{E}_S] = 4(n-4) \Pr[\bar{E}_T | \bar{E}_S] \leq 4n p^3 \approx np^3.$$

$$\Rightarrow \Delta^* \approx n^2 p^5 + np^3 = o(n^4 p^6) = o(|\bar{E}[X]|) \text{ if } n^2 p \gg 1, np \gg 1.$$

Conjecture : $X = \# \text{ of } K_4 \text{ (or other } H\text{)}.$ $|\bar{E}[X]| \rightarrow \infty \Rightarrow X > 0$ whp.

Counterexample. Consider $H = \begin{array}{c} \overset{\circ}{\triangle} \\ \diagdown \quad \diagup \end{array}$ $|\bar{E}[X]| \approx n^5 p^7 \rightarrow \infty \text{ if } p \gg n^{-5/7}$.

However, H contains K_4 . If $p \ll n^{-2/3}$ there is no K_4 in $G(n, p)$.

In fact the threshold function is $n^{-2/3}$ since we can easily extend

a K_4 to H as long as K_4 exists. $n^{-2/3} \gg n^{-5/7}$ why? $\text{Var}[X]$ is

large. no concentration. Δ^* has a term np if two copies of H

overlap with a K_4 and $np \ll n^5 p^7$ if $p \ll n^{-2/3}$. Chebyshov fails.

Definition: The edge-vertex ratio of $G = (V, E)$: $\rho(G) = |E|/|V|$.

The maximum subgraph ratio is given by $m(G) = \max_{H \subseteq G} p(H)$.

Theorem (Bollobás, 1981). Fix a graph $H = (V, E)$. Then $p = n^{-1/m(H)}$

is a threshold function for containing H as a subgraph. Furthermore,

If $p > n^{-1/m(H)}$, then $X_H = \#$ of copies of H in $G(n, p)$ is w.h.p.

$$X_H \approx \mathbb{E}[X] = \binom{n}{|V|} \frac{|V|!}{|\text{Aut}(H)|} p^{|E|} \approx \frac{n^{|V|} p^{|E|}}{|\text{Aut}(H)|}.$$

Proof. Let H' be the subgraph of H achieving the maximum e-v ratio,

i.e. $m(H) = p(H')$. If $p < n^{-1/m(H)}$, $\mathbb{E}[X_{H'}] = o(1) \Rightarrow X_{H'} = 0$ w.h.p.

Now assume $p > n^{-1/m(H)}$. Count the labelled copies of H in $G(n, p)$.

Let L be a labelled copy of H in K_n . A_L be the event of $L \subseteq G(n, p)$.

For fixed L , $\Delta^* = \sum_{L' \sim L} \Pr[A_{L'} | A_L] = \sum_{L' \sim L} p^{|E(L') \setminus E(L)|}$. Note that

of L' that $L' \sim L \approx n^{|V(L') \setminus V(L)|}$, and $p > n^{-1/m(H)} \geq n^{-1/p(L' \cap L)} =$

$n^{-|V(L') \cap V(L)| / |E(L') \cap E(L)|}$. So $\Delta^* \approx \sum n^{|V(L') \setminus V(L)|} p^{|E(L') \setminus E(L)|} \ll n^{|V(L)|} p^{|E(L)|}$

$\Rightarrow \Delta^* \ll \mathbb{E}[X_H] \Rightarrow \text{Var}[X] = \mathbb{E}[X_H]^2 + o(\mathbb{E}[X_H])^2$. \square

Existence of threshold: for which property \mathcal{P} , a threshold exist?

Simpler question: $\Pr[G(n, p) \in \mathcal{P}]$ increasing of p if \mathcal{P} monotone?

Consider a simple case: $\tilde{\mathcal{F}}$ is a family of subset of $[n]$. $\tilde{\mathcal{F}}$ is called

upward closed set (or up-set) if $\forall S \subseteq T$ and $S \in \tilde{F}$, then $T \in \tilde{F}$.

Theorem. Let \tilde{F} be a nontrivial ($\tilde{F} \neq \emptyset, 2^{[n]}$) up-set of $[n]$. Then

$f(p) = \Pr[\text{Bin}(n, p) \in \tilde{F}]$ is a strictly increasing function.

Proof. (by coupling). Coupling: joint distribution of (X, Y) where marginal distributions are given. $0 \leq p < q < 1$. Construct a coupling.

Pick a uniform random vector $(x_1, \dots, x_n) \in [0, 1]^n$. Let $A = \{i : x_i \leq p\}$ and $B = \{j : x_j \leq q\}$. Clearly A has the same distribution as $\text{Bin}(n, p)$ and B has the same distribution as $\text{Bin}(n, q)$.

But $A \subseteq B$. Thus $f(p) = \Pr[A \in \tilde{F}] \leq \Pr[B \in \tilde{F}] = f(q)$. \square

Another proof, also by two-round exposure coupling: Let $0 \leq p < q \leq 1$.

Construct A, B as follows. $\forall i \in [n]$. add i into A with probability

p . If $i \notin A$, add i into B . Otherwise add it into B with probability

$1 - \frac{1-q}{1-p}$. Then $\Pr[i \in B] = p + (1-p)(1 - \frac{1-q}{1-p}) = q$ and independent.

So B has the same distribution as $\text{Bin}(n, q)$, and $A \subseteq B$. \square

Theorem (Bollobás & Thomason, 1987). Every nontrivial monotone (increasing) graph property has a threshold function.

Proof. Consider k independent copies G_1, G_2, \dots, G_k of $G(n, p)$. Their union $G_1 \cup G_2 \cup \dots \cup G_k$ has the same distribution of $G(n, 1 - (1-p)^k)$.

By the monotonicity of \mathcal{P} . If $G_1 \cup G_2 \cup \dots \cup G_k \notin \mathcal{P}$, then $G_1 \cup \dots \cup G_k \notin \mathcal{P}$. By independence, $\Pr[G(n, 1 - (1-p)^k) \notin \mathcal{P}] \leq \Pr[G(n, p) \notin \mathcal{P}]^k$

Let $f(p) = f_n(p) = \Pr[G(n, p) \in \mathcal{P}]$. Note that $(1-p)^k \geq 1 - kp$ by

convexity. A monotone \mathcal{P} , and any positive integer $k \leq 1/p$. we have

$1 - f(kp) \leq 1 - f(1 - (1-p)^k) \leq (1 - f(p))^k$. Fix a sufficiently large n .

Since $f(p)$ is a continuous strictly increasing function from 0 to 1

as p goes from 0 to 1. there is some critical $p_c = p_c(n)$ s.t. $f(p_c) = \frac{1}{2}$.

We claim that p_c is a threshold function. If $p = p(n) \gg p_c$. then

letting $k = kn = \lfloor np/p_c \rfloor \rightarrow \infty$. $1 - f(p) \leq (1 - f(p_c))^k = 2^{-k} \rightarrow 0$. so

$f(p) \rightarrow 1$. Analogously, if $p \ll p_c$. then letting $k = \lfloor np_c/p \rfloor \rightarrow \infty$. we

have $\frac{1}{2} = 1 - f(p_c) \leq (1 - f(p))^k$. Thus $f(p) \rightarrow 0$ as $n \rightarrow \infty$. \square

Remark. In fact, using the method of moments, # of Δ s converges to

a Poisson distribution. Thus $\Pr[\Delta \subseteq G(n, c_n/n)] = \begin{cases} 1 - e^{-c^3/6} & c_n \rightarrow c \\ 1 & c_n \rightarrow \infty \end{cases}$

However, consider some other properties, such as "no isolated v".

We have $\Pr[G(n,p) \text{ has no isolated vertex}] = e^{-e^{-c}}$ if $c_n \rightarrow c$

where $P = \frac{\log n + c_n}{n}$ and $c \in \mathbb{R} \cup \{-\infty, \infty\}$. (exercise!). This

property shows a stronger notion of threshold: sharp threshold.

Definition: We say r_n is a sharp threshold for some graph property \mathcal{P} .

$$\mathcal{P}. \text{ If } \forall \delta > 0, \Pr[G(n, p_n) \in \mathcal{P}] \rightarrow \begin{cases} 0 & \text{if } p_n \leq (1-\delta)r_n \\ 1 & \text{if } p_n \geq (1+\delta)r_n \end{cases}$$

Sharp threshold: $[f^-(\varepsilon), f^+(1-\varepsilon)] = \Theta(r_n) \rightarrow [f^-(\varepsilon), f^+(1-\varepsilon)] = o(r_n)$.

Remark: Roughly speaking, all monotone graph property with a coarse threshold may be approximated by a local property (having some H as a subgraph). (Friedgut's sharp threshold theorem, 1999).

Conjecture: $\forall k \geq 3. \exists d_k$ s.t. $\Pr[G(n, d/n) \text{ is } k\text{-colorable}] \rightarrow \begin{cases} 1 & \text{if } d < d_k \\ 0 & \text{if } d > d_k \end{cases}$

Theorem (Achlioptas & Friedgut, 2000). $\forall k \geq 3. \exists$ function $d_k(n)$ s.t.

$\forall \varepsilon > 0, \Pr[G(n, d(n)/n) \text{ is } k\text{-colorable}] \rightarrow \begin{cases} 1 & \text{if } d(n) < d_k(n) - \varepsilon \\ 0 & \text{if } d(n) > d_k(n) + \varepsilon \end{cases}$

Open question: whether $d_k(n)$ has a limit d_k ?

Consider an easy case: chromatic number of $G(n, 1/2)$. We do

know it has concentration in the course Advanced Algorithms.

Start from a question: what is the clique number of $G(n, 1/2)$?

Let $X = \#$ of cliques of $G(n, 1/2)$. We have $f_k = \mathbb{E}[X] = \binom{n}{k} 2^{-\binom{k}{2}}$.

Fix S of size k . $\forall T$ of size k . $T \sim S \Rightarrow |S \cap T| \geq 2$. So we have

$$\begin{aligned}\Delta^* &= \sum_{T \sim S} \Pr[A_T \mid A_S] = \sum_{i=2}^{k-1} \binom{k}{i} \binom{n-k}{k-i} 2^{\binom{i}{2} - \binom{k}{2}} \leq 2^{-\binom{k}{2}} \sum k^i n^{k-i} 2^{\binom{i}{2}} \\ &\leq 2^{-\binom{k}{2}} n^k \sum \left(\frac{k 2^i}{n}\right)^i = o(f_k) \text{ if } f_k \rightarrow \infty. \Rightarrow \Pr[X > 0] = 1 - o(1).\end{aligned}$$

Theorem. (Bollobás - Erdős, 1976. Matula 1976) $\omega(G(n, 1/2)) \approx 2 \log n$.

Corollary: $\alpha(G(n, 1/2)) \approx 2 \log_2 n \Rightarrow \chi(G(n, 1/2)) \geq (1 + o(1)) \frac{n}{2 \log_2 n}$.